

IE605: Engineering Statistics

Lecture 13

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Previous Lecture:

- ▶ Evaluating Estimators
- ▶ Cramer's Rao inequality

This Lecture:

- ▶ Cramer's Rao inequality
- ▶ Information inequality
- ▶ Fischer Information
- ▶ Hypothesis testing
- ▶ Likelihood Ratio Test (LRT)

Cramer-Rao's Bound

Let $\mathbf{X} = (X_1, X_2, \dots, X_n)$ be a sample with pdf $f(\mathbf{x}|\theta)$, and let $W(\mathbf{X})$ be any estimator such that $\text{Var}_\theta W(\mathbf{X}) < \infty$. Then

$$\text{Var}_\theta W(\mathbf{X}) \geq \frac{\left(\frac{d}{d\theta} \mathbb{E}_\theta W(\mathbf{X})\right)^2}{\mathbb{E}_\theta \left(\left(\frac{\partial}{\partial \theta} \log f(\mathbf{X}|\theta)\right)^2 \right)}$$

provided

$$\mathbb{E}_\theta W(\mathbf{X}) = \int_{\mathbf{x}} \frac{\partial}{\partial \theta} [W(\mathbf{x})f(\mathbf{x})] d\mathbf{x}.$$

Proof of Cramer-Rao's Bound

▶ $[\text{Cov}(X, Y)]^2 \leq \text{Var}(X)\text{Var}(Y) \implies \text{Var}(X) \geq \frac{[\text{Cov}(X, Y)]^2}{\text{Var}(Y)}$



$$\frac{d}{d\theta} \mathbb{E}_\theta W(\mathbf{X}) = \int_{\mathbf{x}} \frac{\partial}{\partial \theta} [W(\mathbf{x})f(\mathbf{x})] d\mathbf{x}$$



$$\mathbb{E}_\theta \left(\frac{\partial}{\partial \theta} \log f(\mathbf{X}|\theta) \right)$$



$$\text{Cov} \left(W(\mathbf{X}), \frac{\partial}{\partial \theta} \log f(\mathbf{X}|\theta) \right) = \mathbb{E}_\theta \left(W(\mathbf{X}) \frac{\partial}{\partial \theta} \log f(\mathbf{X}|\theta) \right)$$

Cramer-Rao Bound for iid case

$$\begin{aligned} \text{Var}_\theta W(\mathbf{X}) &\geq \frac{\left(\frac{d}{d\theta} \mathbb{E}_\theta W(\mathbf{X})\right)^2}{\mathbb{E}_\theta \left(\left(\frac{\partial}{\partial \theta} \log f(\mathbf{X}|\theta)\right)^2 \right)} \\ &= \frac{\left(\frac{d}{d\theta} \mathbb{E}_\theta W(\mathbf{X})\right)^2}{n \mathbb{E}_\theta \left(\left(\frac{\partial}{\partial \theta} \log f(X|\theta)\right)^2 \right)} \end{aligned}$$

$$\mathbb{E}_\theta \left(\left(\frac{\partial}{\partial \theta} \log f(\mathbf{X}|\theta)\right)^2 \right) = \mathbb{E}_\theta \left(\left(\sum_i \frac{\partial}{\partial \theta} \log f(X_i|\theta)\right)^2 \right)$$

Fisher Information and Information Inequality

The quantity $\mathbb{E}_\theta \left(\left(\frac{\partial}{\partial \theta} \log f(\mathbf{X}|\theta) \right)^2 \right)$ is called **Information number** or **Fisher Information of sample**

- ▶ Larger the information number, smaller is the bound in the Cramer-Rao's bound
- ▶ Larger the information number, we have more information about about θ .
- ▶ Cramer-Rao bound is also called as Information inequality.

Hypothesis Testing

Definition: A Hypothesis is a statement about a population parameter

Definition: Two complementary hypothesis in a hypothesis testing are called *null hypothesis* and *alternate hypothesis*, denoted as H_0 and H_1 , respectively.

General form of hypothesis testing

$$H_0 : \theta \in \Theta_0 \quad \text{null hypothesis}$$

$$H_1 : \theta \in \Theta_0^c \quad \text{null hypothesis}$$

$$H_0 : \theta \leq \theta_0$$

$$H_1 : \theta > \theta_0$$

for some threshold θ_0 .

Hypothesis Testing contd..

Hypothesis testing procedure or hypothesis test is rule that prescribes

1. For which sample values the **decision** is made to accept H_0 as true
2. For which sample values H_0 is rejected and H_1 is accepted as true

Methods of tests: Likelihood Ratio Test (LRT)

Likelihood ratio test statistic for testing $H_0 : \theta \in \Theta_0$ and $H_1 : \theta \in \Theta_0^c$ is

$$\lambda(\mathbf{x}) = \frac{\sup_{\theta \in \Theta_0} L(\theta|\mathbf{x})}{\sup_{\theta \in \{\Theta_0, \Theta_0^c\}} L(\theta|\mathbf{x})}$$

A Likelihood Ratio Test (LRT) is any test that has the reject region of the form $\{\mathbf{x} : \lambda(\mathbf{x}) \leq c\}$ for some $c \in (0, 1)$.

Example:

Sufficient Statistics and LRT

Let $T(\mathbf{X})$ be a sufficient statistic of sample $\mathbf{X} = \{X_1, X_2, \dots, X_n\}$ drawn from population $f(\mathbf{x}|\theta)$

$$\begin{aligned}\lambda(\mathbf{x}) &= \frac{\sup_{\theta \in \Theta_0} L(\theta|\mathbf{x})}{\sup_{\theta \in \Theta} L(\theta|\mathbf{x})} \\ &= \frac{\sup_{\theta \in \Theta_0} f(\mathbf{x}|\theta)}{\sup_{\theta \in \Theta} f(\mathbf{x}|\theta)} \\ &= \frac{\sup_{\theta \in \Theta_0} h(\mathbf{x})g(T(\mathbf{x})|\theta)}{\sup_{\theta \in \Theta} h(\mathbf{x})g(T(\mathbf{x})|\theta)} \\ &= \frac{\sup_{\theta \in \Theta_0} g(T(\mathbf{x})|\theta)}{\sup_{\theta \in \Theta} g(T(\mathbf{x})|\theta)} \\ &= \frac{\sup_{\theta \in \Theta_0} L^*(\theta|T(\mathbf{x}))}{\sup_{\theta \in \Theta} L^*(\theta|T(\mathbf{x}))} = \lambda(T(\mathbf{x}))\end{aligned}$$

Bayes Tests:

In Baye's method, we have prior probability $P(\theta)$ and posterior probability $\pi(\theta|\mathbf{x})$ after observing sample \mathbf{x} .

- ▶ $\pi(\theta|\mathbf{x})$ can be used to find

$$\Pr(\theta \in \Theta_0|\mathbf{x}) \quad \text{and} \quad \Pr(\theta \in \Theta_0^c|\mathbf{x})$$

- ▶ We can assign

$$\Pr(\theta \in \Theta_0|\mathbf{x}) = \Pr(H_0 \text{ is true} | \mathbf{x})$$

$$\Pr(\theta \in \Theta_0^c|\mathbf{x}) = \Pr(H_1 \text{ is true} | \mathbf{x})$$

- ▶ Decision criteria

H_0 is true if $\Pr(\theta \in \Theta_0|\mathbf{x}) \geq \Pr(\theta \in \Theta_0^c|\mathbf{x})$

H_1 is true otherwise

Rejection set $\{\mathbf{x} : \Pr(\theta \in \Theta_0^c|\mathbf{x}) \geq 1/2\}$.