

**IE 605: Engineering Statistics**

Solutions: Tutorial 1

**Solution 1**

$$\Omega = \{\{i, j, k\} \mid i, j, k \in \{1, 2, 3, 4, 5, 6\}\}$$

$$F = 2^\Omega$$

$$P(\text{Same no. on exactly 2 dice}) = \frac{\binom{3}{2} \times 6 \times 5}{6^3} = 5/12.$$

**Solution 2**

1.  $A \cup B \cup C$
2.  $(A \cup B \cup C)^c \cup (A \cup B)^c \cup (A \cup C)^c \cup (C \cup B)^c$
3.  $(A \cup B \cup C)^c$
4.  $A \cap B \cap C$
5.  $(A \cap B^c \cap C^c) \cup (A^c \cap B \cap C^c) \cup (A^c \cap B^c \cap C)$
6.  $A \cap B \cap C^c$
7.  $A \cup (A \cup B)^c$

**Solution 3**

$$\Omega = \{HH, THH, HTHH, TTHH, \dots\}$$

There are two configurations for exactly 4 tosses i.e.  $\{TTHH, HTHH\}$

$$P(\text{exactly 4 tosses}) = \left(\frac{1}{2}\right)^4 + \left(\frac{1}{2}\right)^4 = \frac{1}{8}$$

**Solution 4**

$$n(\text{geniuses}) = 60$$

$$n(\text{chocolate lovers}) = 70$$

$$n(\text{geniuses} \cap \text{chocolate lovers}) = 40$$

$$\text{So that, } n(\text{geniuses} \cup \text{chocolate lovers}) = 70 + 60 - 40 = 90$$

$$P(\text{geniuses} \cup \text{chocolate lovers})^c = \frac{100-90}{100} = 0.10$$

## Solution 5

Let  $P(\text{odd face}) = 2p$ ;

$P(\text{even face}) = p$ ;

As probability of sample space will be 1;

so,

$$P(1) + P(2) + P(3) + P(4) + P(5) + P(6) = 1$$

$$2p + p + 2p + p + 2p + p = 1$$

$$p = \frac{1}{9}$$

Now,

$$P(1) + P(2) + P(3) = 4p = \frac{4}{9}$$

## Solution 6

1. Given  $A \subset B$ ;

Let  $D = B/A$  or  $B \cap A^c$ . then A and D will be mutually exclusive sets;

hence,  $P(B) = P(D) + P(A)$ ;

so,  $P(B) \geq P(A)$

2. Let  $D_1 = A$

$$D_2 = B \cap A^c$$

$A \cup B = D_1 \cup D_2$ , where  $D_1$  and  $D_2$  are mutually exclusives.

$$P(A \cup B) = P(D_1) + P(D_2)$$

$$\implies P(A \cup B) = P(A) + P(B \cap A^c) \quad \dots\dots\dots(i)$$

Again,  $B = (B \cap A^c) \cup (A \cap B)$ , where  $B \cap A^c$  and  $A \cap B$  are mutually exclusives.

$$\implies P(B) = P(B \cap A^c) + P(A \cap B) \quad \dots\dots\dots(ii)$$

Inserting (ii) in (i) we get,  $P(A \cup B) = P(A) + P(B) - P(A \cap B)$ .

3.

$$P(A \cup B \cup C) = P((A \cup B) \cup C)$$

$$= P(A \cup B) + P(C) - P((A \cup B) \cap C)$$

$$= P(A) + P(B) + P(C) - P(A \cap B) - P((A \cap C) \cup (B \cap C))$$

$$= P(A) + P(B) + P(C) - P(A \cap B) - [P(A \cap C) + P(B \cap C) - P(A \cap B \cap C)]$$

Therefore,

$$P(A \cup B \cup C) = P(A) + P(B) + P(C) - P(A \cap B) - P(A \cap C) - P(B \cap C) + P(A \cap B \cap C).$$

$$4. \Omega = A \cup A^c$$

As A and  $A^c$  are mutually exclusive

$$P(A \cup A^c) = P(A) + P(A^c)$$

$$P(\Omega) = P(A) + P(A^c);$$

$$P(A) + P(A^c) = 1$$

5. As  $\Omega$  and  $\phi$  are mutually exclusive;

$$P(\Omega \cup \phi) = P(\Omega) + P(\phi);$$

$$P(\phi) = 0$$

## Solution 7

**Proof for Axiom 1:**

$$P'(A) = P(A|B) = \frac{P(A \cap B)}{P(B)}.$$

As  $P(A \cap B) \geq 0$

hence  $P'(A) \geq 0$

**Proof for Axiom 2:**

$$P'(\Omega) = P(\Omega|B) = \frac{P(\Omega \cap B)}{P(B)} = \frac{P(B)}{P(B)} = 1.$$

**Proof for Axiom 3:**

Let A and B are two mutually exclusive set;

$$P'(A \cup C) = P((A \cup C)|B);$$

$$P'(A \cup C) = \frac{P((A \cup C) \cap B)}{P(B)}$$

$$P'(A \cup C) = \frac{P((A \cap B) \cup (C \cap B))}{P(B)}$$

$$P'(A \cup C) = \frac{P(A \cap B)}{P(B)} + \frac{P(C \cap B)}{P(B)}$$

$$P'(A \cup C) = P(A|B) + P(C|B)$$

$$P'(A \cup C) = P'(A) + P'(C)$$

## Solution 8

Let, we toss two coin simultaneously,

we are assuming the following events;

A = both coin agree ( HH, TT );

B = head appears on coin 1 ( HH, HT );

C = head appears on coin 2 ( HH , TH );

For independence of two sets A and B ;

$$P(A \cap B) = P(A).P(B)$$

we can see,

$$P(A \cap B) = P(A) \cdot P(B) = \frac{1}{4}$$

$$P(A \cap c) = P(A) \cdot P(C) = \frac{1}{4}$$

$$P(C \cap B) = P(C) \cdot P(B) = \frac{1}{4}$$

$$\text{Now, For } P(A \cap B \cap C) = \frac{1}{4}$$

$$\text{but } P(A) \cdot P(B) \cdot P(C) = \frac{1}{8}$$

hence pairwise Independence does not make mutual independence.

## Solution 9

Given  $A_i \in \mathcal{F}, i \geq 1$ .

$$\text{To show: } P(\cup_{i=1}^{\infty} A_i) \leq \sum_{i=1}^{\infty} P(A_i)$$

Consider  $B_i = A_i - \cup_{j=1}^{i-1} A_j$ , such that  $\{B_i\}_{i \geq 1}$  are disjoint.

By construction,  $B_i \subset A_i$  and  $\cup_{i=1}^{\infty} B_i = \cup_{i=1}^{\infty} A_i$

Using Ques 6(1),  $P(B_i) \leq P(A_i)$

$$\text{So, } P(\cup_{i=1}^{\infty} A_i) = P(\cup_{i=1}^{\infty} B_i) = \sum_{i=1}^{\infty} P(B_i) \leq \sum_{i=1}^{\infty} P(A_i).$$

## Solution 10

Let A be the event that the product is manufactured at company A, B be the event that the product is manufactured at company B, D be the event that the product is defective.

$$1. P(\text{sample is defective}) = P(A)P(D|A) + P(B)P(D|B) = 0.26$$

$$2. P(\text{defective sample is manufactured at company A}) = P(A|D) = \frac{P(A \cap D)}{P(D)} = \frac{12}{13}$$

## Solution 11

Let F denote the event that the attempt fails, W be the event that the server is working, NW be the event that the server is not working.

$$1. P(\text{first attempt fails}) = P(F|NW)P(NW) + P(F|W)P(W) = 0.28.$$

$$2. P(\text{server was working given that first access attempt fails}) = \frac{P(F \cap W)}{P(F)} = \frac{P(F|W)P(W)}{P(F)} = 2/7.$$

3. Let  $F_1, F_2$  denote the event that the first, second access attempt fails respectively.

$$P(F_2|F_1) = \frac{P(F_1, F_2)}{P(F_1)} = \frac{P(W)P(F_1|W)P(F_2|W) + P(NW)P(F_1|NW)P(F_2|NW)}{P(F_1)} = 208/280.$$

$$4. P(\text{server is working given that first and second attempt fails}) = \frac{P(W \cap F_1 \cap F_2)}{P(F_1 \cap F_2)} = \frac{P(F_1 \cap F_2|W)P(W)}{P(F_1 \cap F_2)} = 8/208.$$

## Solution 12

We will use first principle of mathematical induction for this question.

For  $n = 2$ , refer to ques 6. Let the claim be true for  $n = r$ . We will thus prove that it also holds for  $n = r + 1$ .

$$\begin{aligned}
 P(\cup_{i=1}^{r+1} E_i) &= P((\cup_{i=1}^r E_i) \cup E_{r+1}) \\
 &= P(\cup_{i=1}^r E_i) + P(E_{r+1}) - P((\cup_{i=1}^r E_i) \cap E_{r+1}) \\
 &= P(\cup_{i=1}^r E_i) + P(E_{r+1}) - P(\cup_{i=1}^r (E_i \cap E_{r+1})) \\
 &= \sum_{i=1}^r P(E_i) - \sum_{1 \leq i < j \leq r} P(E_i \cap E_j) + \cdots + (-1)^{r+1} P(\cap_{i=1}^r E_i) + P(E_{r+1}) \\
 &\quad - P(\cup_{i=1}^r (E_i \cap E_{r+1})) \\
 &= \sum_{i=1}^{r+1} P(E_i) - \sum_{1 \leq i < j \leq r} P(E_i \cap E_j) + \cdots + (-1)^{r+1} P(\cap_{i=1}^r E_i) - \left\{ \sum_{i=1}^r P(E_i \cup E_{r+1}) \right\} \\
 &\quad - \sum_{1 \leq i < j \leq r} P((E_i \cap E_{r+1}) \cap (E_j \cap E_{r+1})) + \cdots + (-1)^{r+1} P(\cap_{i=1}^{r+1} E_i) \} \\
 &= \sum_{i=1}^{r+1} P(E_i) - \left\{ \sum_{1 \leq i < j \leq r} P(E_i \cap E_j) + \sum_{i=1}^r P(E_i \cap E_{r+1}) \right\} + \cdots + (-1)^{r+2} P(\cap_{i=1}^{r+1} E_i) \\
 &= \sum_{i=1}^{r+1} P(E_i) - \sum_{1 \leq i < j \leq r+1} P(E_i \cap E_j) + \cdots + (-1)^{r+2} P(\cap_{i=1}^{r+1} E_i)
 \end{aligned}$$

Thus, the claim holds true for any finite value of  $n$ .

## Solution 13

Let  $B$  = event both are girls;  $E$  = event oldest is girl and  $L$  = event at least one is a girl.

Part (a):

$$P(B|E) = \frac{P(BE)}{P(E)} = \frac{P(B)}{P(E)} = \frac{1/4}{1/2} = \frac{1}{2}.$$

Part (b):

$$P(L) = 1 - P(\text{no girls}) = 1 - \frac{1}{4} = \frac{3}{4}$$

Then,

$$P(B|L) = \frac{P(BL)}{P(L)} = \frac{P(B)}{P(L)} = \frac{1/4}{3/4} = \frac{1}{3}$$

## Solution 14

Let  $E$  = event at least one is a six.

$$P(E) = \frac{\text{number of ways to get E}}{\text{number of sample points}} = \frac{11}{36}$$

Let  $D$  = event two faces are different

$$P(D) = 1 - \text{Prob}(\text{two faces the same}) = 1 - \frac{6}{36} = \frac{5}{6}$$

Then,

$$P(E|D) = \frac{P(ED)}{P(D)} = \frac{10/36}{5/6} = \frac{1}{3}$$

## Solution 15

Let  $E$  = event same number on exactly two of the dice;  $S$  = event all three numbers are the same;  $D$  = event all three numbers are different. These three events are mutually exclusive and define the whole sample space. Thus,

$$1 = P(D) + P(S) + P(E)$$

Now,

$$P(S) = \frac{6}{216} = \frac{1}{36}$$

For even  $D$ , we have six possible values for first die, five for second, and four for third. Therefore, Number of ways to get  $D = 6 \cdot 5 \cdot 4 = 120$ .

$$\implies P(D) = \frac{120}{216} = \frac{20}{36}$$

Then,

$$\begin{aligned} P(E) &= 1 - P(S) - P(D) \\ &= 1 - \frac{1}{36} - \frac{20}{36} \\ &= \frac{5}{12}. \end{aligned}$$

## Solution 16

Let  $C$  = event person is color blind. Then,

$$\begin{aligned} P(\text{Male}|C) &= \frac{P(C|\text{Male})P(\text{Male})}{P(C|\text{Male})P(\text{Male}) + P(C|\text{Female})P(\text{Female})} \\ &= \frac{0.05 \times 0.5}{0.05 \times 0.5 + 0.0025 \times 0.5} \end{aligned}$$

$$= \frac{2500}{2625} = \frac{20}{21}.$$

## Solution 17

Let  $A_k$ : points of A in game  $k$ .

$B_k$ : points of B in game  $k$ .

Let  $x_k = A_k - B_k$ : difference of points between A and B in game  $k$ .

Given:  $\mathbb{P}\{A_k\} = p \forall k = 1, 2, \dots$

$\mathbb{P}\{\text{Total } 2n \text{ points}\} = \mathbb{P}\{X_{2n} \in \{-2, 2\}, X_{2n-2} = 0, X_{2n-4} = 0, \dots, X_0 = 0\}$

$$\mathbb{P}\{X_0 = 0\} = 1$$

$$\mathbb{P}\{X_2 = 0, X_0 = 0\} = p(1-p) + (1-p)p = 2p(1-p)$$

$$\begin{aligned} \mathbb{P}\{X_4 = 0, X_2 = 0, X_0 = 0\} &= p(1-p)p(1-p) + p(1-p)(1-p)p \\ &\quad + (1-p)pp(1-p)(1-p)p(1-p)p \\ &= 2^2 p^2 (1-p)^2 \end{aligned}$$

$X_0 = 0; X_1 = -1$  (if B wins);  $X_2 = 0$  (if A wins);  $X_3 = 1$  (if A wins);  $X_4 = 0$  (if B wins)  $\implies (1-p)pp(1-p)$  holds.

Similarly,  $X_0 = 0; X_1 = 1$  (if A wins);  $X_2 = 0$  (if B wins);  $X_3 = -1$  (if B wins);  $X_4 = 0$  (if A wins)  $\implies p(1-p)(1-p)p$  holds.

Similarly,

$$\mathbb{P}\{X_{2n-2} = 0, \dots, X_0 = 0\} = 2^{n-1} p^{n-1} (1-p)^{n-1}$$

Therefore,

$$\begin{aligned} \mathbb{P}\{X_{2n} \in \{-2, 2\}, X_{2n-2} = 0, X_{2n-4} = 0, \dots, X_0 = 0\} &= 2^{n-1} p^{n-1} (1-p)^{n-1} (p^2 + (1-p)^2) \\ \mathbb{P}\{\text{A wins}\} &= \sum_{n=1}^{\infty} \{2^{n-1} p^{n-1} (1-p)^{n-1} p^2\} = \frac{p^2}{1 - 2p(1-p)}. \end{aligned}$$